

# KAPLAN-MEIER ESTIMATE

Dr. Mutua Kilai

Department of Pure and Applied Sciences

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# Kaplan-Meier Estimator of Survival

- Let  $t_i$  denote an ordered observed value. The **empirical survivor function(esf)** denoted by  $S_n(t)$  is defined by:

$$S_n(t) = \frac{\text{Number of observations} > t}{n} = \frac{\{t_i > t\}}{n} \quad (1)$$

- Example: Calculate the Empirical Survivor function for the following data:

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9	13	13	18	23	28	31	34	45	48	161
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# Solution

t	0	9	13	18	23	28	31	34	45	48	161
$S_n(t)$	$\frac{11}{11}$	$\frac{10}{11}$	$\frac{8}{11}$	$\frac{7}{11}$	$\frac{6}{11}$	$\frac{5}{11}$	$\frac{4}{11}$	$\frac{3}{11}$	$\frac{2}{11}$	$\frac{1}{11}$	0

```
library(survival)
```

```
aml <- c(9,13,13,18,23,28,31,34,45,48,16)
```

```
status <- rep(1,11)
```

```
esf.fit <- survfit(Surv(aml, status)~1)
```

```
plot(esf.fit,conf.int=F,xlab="time until relapse (in weeks)"  
ylab="proportion without relapse",lab=c(10,10,7))
```

- The **esf** is a consistent estimator of the true survivor function  $S(t)$ .
- The exact distribution of  $nS_n(t)$  for each fixed  $t$  is binomial  $(n, p)$  where  $n$  = number of observations and  $p = P(T > t)$

# Kaplan-Meier Estimate

- The K-M adjusts the esf in order to reflect the presence of right-censored observations.
- Let  $y_{(i)}$  denote the  $i^{th}$  distinct ordered censored or uncensored observation and is the right interval of the interval  $I_i$
- Let  $R(t)$  denote the **risk set just before time t** and let:
  - $n_i$  denote the number in  $R(y_{(i)})$  and the number alive and not censored just before  $y_{(i)}$
  - $d_i$  the number died at time  $y_{(i)}$
  - $p_i$   $P(\text{surviving through } I_i | \text{alive at beginning } I_i)$
  - $q_i = 1 - p_i = P(\text{die in } I_i | \text{alive at beginning } I_i)$
- Recall the general multiplication of joint events  $A_1$  and  $A_2$

$$P(A_1 \cap A_2) = P(A_2|A_1)P(A_1)$$

## Cont'd

- From this repeated rule the survivor function can be expressed as:

$$S(t) = P(T > t) = \prod_{y_{(i)} \leq t} p_i$$

- The estimates of  $p_i$  and  $q_i$  are:

$$\hat{q}_i = \frac{d_i}{n_i}$$

- and

$$\hat{p}_i = 1 - \hat{q}_i = 1 - \frac{d_i}{n_i}$$

- Formally

$$\hat{S}(t) = \prod_{y_{(i)} \leq t} \hat{p}_i = \prod_{y_{(i)} \leq t} \left( \frac{n_i - d_i}{n_i} \right)$$

where

- $n_i$  the number of subjects at risk at time  $t_i$
- $d_i$  is the number of individuals who fail at time  $t_i$

# Advantages of K-M Estimate

- It is simple and straightforward to use and interpret
- it is a nonparametric estimator, so it constructs a survival curve from the data and no assumptions is made about the shape of the underlying distribution
- it gives a graphical representation of the survival function(s), useful for illustrative purposes



# Example

Consider the following data and calculate the K-M estimate

subject	time	event
1	3	0
2	5	1
3	7	1
4	2	1
5	18	0
6	16	1
7	2	1
8	9	1
9	16	1
10	5	0

where

- subject: is the individuals identifier
- time: is the time to event (in years)
- event: is the event status (0 = censored, 1 = event happened)

# Solution

- We first need to count the number of distinct event times, ignoring censored observations we have 5 distinct event times.
- We make a table and fill as follows;
  - $y_{(j)}$  gives the ordered distinct event times
  - $d_j$  gives the number of observations for each distinct event time
  - $R_j$  gives the remaining number of individuals at risk. For this, the distribution of time (censored and not censored) is useful.

$y_j$	$d_j$	$R_j$	$1 - \frac{d_j}{R_j}$	$S_{KM}(t)$
2	2	10	0.800	0.800
5	1	7	0.857	0.686
7	1	5	0.800	0.548
9	1	4	0.750	0.411
16	2	3	0.333	0.317

- We can compute the KM estimator in R using the following.

We enter the data in R

```
# create dataset  
dat <- data.frame(  
  time = c(3, 5, 7, 2, 18, 16, 2, 9, 16, 5),  
  event = c(0, 1, 1, 1, 0, 1, 1, 1, 1, 0))
```

We then run the K-M estimator using the `survfit()` and `Surv()` functions as follows:

```
# KM  
library(survival)  
  
km <- survfit(Surv(time, event) ~ 1,  
  data = dat  
)
```

- The `Surv()` function accepts two arguments:
  - the time variable
  - the event variable
- The `~` in the `survfit()` function indicates that we estimate the Kaplan-Meier without any grouping.
- We can plot the K-M as follows

```
library(survminer)

# plot
ggsurvplot(km,
  conf.int = FALSE,
  legend = "none"
)
```

- The crosses on the survival curve denote the censored observations.
- The advantage with the `ggsurvplot()` function is that it is easy to draw the median survival directly on the plot:

```
ggsurvplot(km,  
  conf.int = FALSE,  
  surv.median.line = "hv",  
  legend = "none"  
)
```

# Confidence Interval

- To obtain the confidence limits for the product limit estimator, we first use the delta method in order to obtain the variance of  $\log(\hat{S}(t))$
- The delta method allows one to approximate the variance of a continuous variable  $g(\cdot)$  of a random variable.
- If a random variable  $X$  has mean  $\mu$  and variance  $\sigma^2$  then  $g(X)$  will have approximate mean  $g(\mu)$  and variance  $\sigma^2 \times [g'(\mu)]^2$  for a sufficiently large sample size.
- Applying the delta's formula we get the variance of  $\log \hat{S}(t)$  as

$$\text{var}\left(\log \hat{S}(t_k)\right) = \sum_{t_i \leq t} \text{var} \log(1 - \hat{q}_i) \approx \sum_{t_i \leq t} \frac{d_j}{n_j(n_j - d_j)} \quad (2)$$



- To get the variance of  $\hat{S}_t$  itself we use the delta method again to obtain:

$$\text{var}\left(\hat{S}(t)\right) \approx [\hat{S}(t)]^2 \sum_{t_i \leq t} \frac{d_i}{n_i(n_i - d_i)} \quad (3)$$

- Unfortunately, confidence intervals computed based on this variance may extend above one or below zero.
- A more satisfying approach is to find the confidence intervals for the complementary log-log transformation of  $\hat{S}(t)$  as follows:

$$\text{var}\left(\log \left[ -\log \hat{S}(t) \right]\right) \approx \frac{1}{[\log \hat{S}(t)]^2} \sum_{t_i \leq t} \frac{d_i}{n_i(n_i - d_i)} \quad (4)$$

- Theory tells us that for each fixed value  $t$

$$\hat{S}(t) \sim N(S(t), \text{var}\left(\hat{S}(t)\right))$$

- Thus at time  $t$  an approximate  $(1 - \alpha) \times 100\%$  confidence interval for the probability of survival  $S(t) = P(T > t)$  is given by:

$$\hat{S}(t) \pm z_{\frac{\alpha}{2}} s.e(\hat{S}(t))$$

# Exercise

Find the K-M estimator for the following data ( $n = 21$ ) and obtain the 95% CI for  $S(t)$  when  $t = 21$

6, 6, 6, 6<sup>+</sup>, 7, 9<sup>+</sup>, 10, 10<sup>+</sup>, 11<sup>+</sup>, 13, 16, 17<sup>+</sup>, 19<sup>+</sup>, 20<sup>+</sup>, 22, 23, 25<sup>+</sup>, 32<sup>+</sup>, 32<sup>+</sup>

Thank You!